



University
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MULTIVARIATE FORECASTING IN ENERGY SYSTEMS WITH A LARGE SHARE OF RENEWABLES

PIMS Workshop on Forecasting and Mathematical Modeling for Renewable Energy

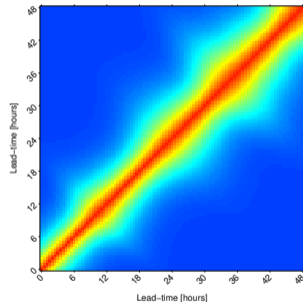
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MOTIVATION

Decision-makers (TSOs, DSOs, traders) require forecasts of multiple quantities to operate efficiently and manage risk:

- How much demand will be met by wind and solar power tomorrow?
- What is the chance of power flows exceeding network capacity?
- Will I get a better price if I sell my power day-ahead or intraday?

Probabilistic forecasts quantify uncertainty by expressing predictions as probability density functions.

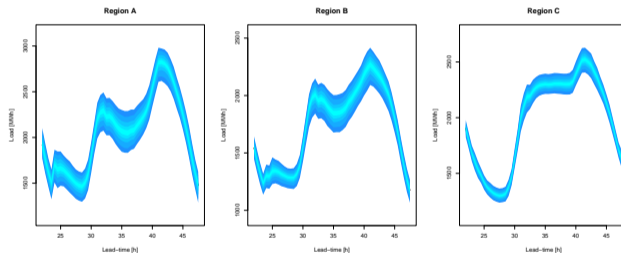


Figure 1: Fan plots of density forecasts for three locations and 48 time periods

⚠ These forecasts do not describe spatial or temporal dependency.

Probabilistic forecasts quantify uncertainty by expressing predictions as probability density functions.

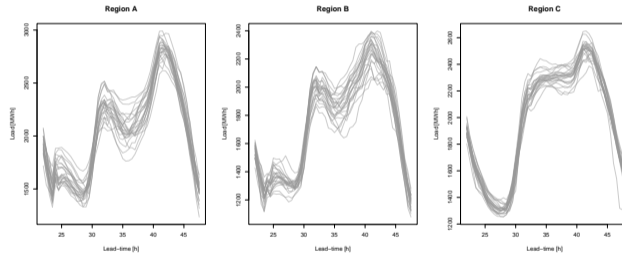


Figure 2: Space-time trajectories (or scenarios/samples) drawn from multivariate probabilistic forecast. These forecasts contain dependency information but are difficult to visualise.

This quickly becomes a **high-dimensional** problem!

Gaussian copulas provide a suitable framework for describing such high-dimensional predictive distributions:

- Margins of the copula are the familiar density forecasts
- Dependency structure specified by a covariance matrix, Σ
- Scales well (compared to other copulas) but limited by estimation of the covariance matrix

The remainder of this talk is concerned with this **covariance matrix** and the possibility that it:

1. has a complex structure, and/or
2. varies over time, perhaps as a function of some covariate.

COVARIANCE FUNCTIONS AND MATRICES

Consider a random process $Z_t(\mathbf{s}, l)$ at location \mathbf{s} , forecast lead-time l , and forecast issue time t .

A covariance function, C_t , is *stationary* if the covariance

$$\text{cov}(Z_t(\mathbf{s}, l), Z_t(\mathbf{s} + \mathbf{h}, l + u)) = C_t(\mathbf{h}, u) \quad (1)$$

depends only on separation (\mathbf{h}, u) . Furthermore, C_t is *isotropic* if it is invariant to the direction of \mathbf{h} and u

$$\text{cov}(Z_t(\mathbf{s}, l), Z_t(\mathbf{s} + \mathbf{h}, l + u)) = C_t(\|\mathbf{h}\|, |u|) \quad (2)$$

Table 1: Some parametric classes of isotropic covariance functions where $C(\mathbf{h})$ takes the form $C(\|\mathbf{h}\|; \boldsymbol{\xi})$. The Whittle–Matérn covariance is defined in terms of the modified Bessel function of the second kind K_ν .

Class	Function $C(r; \boldsymbol{\xi})$	Parameters $\boldsymbol{\xi}$
Powered Exponential	$\sigma^2 e^{-(\theta r)^\gamma}$	$0 < \gamma \leq 2; \theta > 0; \sigma \geq 0$
Whittle–Matérn	$\sigma^2 \frac{2^{1-\nu}}{\Gamma(\nu)} (\theta r) K_\nu(\theta r)$	$\nu > 0; \theta > 0; \sigma \geq 0$
Cauchy	$\sigma^2 (1 + (\theta r)^\gamma)^{-\nu}$	$0 < \gamma \leq 2; \nu > 0; \theta > 0; \sigma \geq 0$
Spherical	$\sigma^2 \left(1 - \frac{2}{\pi} \left(\frac{r}{\theta} \sqrt{1 - \left(\frac{r}{\theta} \right)^2} + \sin^{-1} \frac{r}{\theta} \right) \right)$	$c(r) = 0$ if $r > \theta$; $\sigma^2 \geq 0; \theta > 0$

Given the separation, $(\|\mathbf{h}\|, |u|)$ between all pairs of variables i and j , given by $R_{i,j}$, the dynamic (time-dependent) covariance matrix Σ_t may be formed as

$$\Sigma_t = \begin{pmatrix} C_t(R_{1,1}) & C_t(R_{1,2}) & \dots & C_t(R_{1,p}) \\ C_t(R_{2,1}) & \ddots & & \vdots \\ \vdots & & & \\ C_t(R_{p,1}) & \dots & & C_t(R_{p,p}) \end{pmatrix} . \quad (3)$$

Therefore, we can specify a covariance matrix of arbitrary size by a covariance function, **with a small number of parameters**, and the known separation matrix, R .

Alternatively we can consider a matrix decomposition of the covariance matrix or its inverse, in this case the Modified Cholesky Decomposition of Σ^{-1} [Pourahmadi, 1999],

$$\Sigma^{-1} = \mathbf{T}^T \mathbf{D}^{-2} \mathbf{T}, \quad (4)$$

where \mathbf{D}^2 is a diagonal matrix with $D_{jj}^2 = \exp(\eta_{j+d})$, for $j = 1, \dots, d$, and

$$\mathbf{T} = \begin{pmatrix} 1 & 0 & 0 & \cdots & 0 \\ \eta_{2d+1} & 1 & 0 & \cdots & 0 \\ \eta_{2d+2} & \eta_{2d+3} & 1 & \cdots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ \eta_{q-d+2} & \eta_{q-d+3} & \cdots & \eta_q & 1 \end{pmatrix}. \quad (5)$$

However, what if the dependency structure we would like to model is

- Non-stationary, i.e. $C_t(\cdot)$ depends on specific location \mathbf{s} or lead time l , or
- Dynamic, evolves over time or via a random process or via dependence on a time-varying covariate?

We can model these behaviours in a parsimonious fashion by allowing the parameters to covariance functions $C_t(\cdot)$, or elements of the modified Cholesky decomposition η_i to be additive models of covariates (which may include \mathbf{s} and/or l).

FLEXIBLE COVARIANCE MODELLING

Let $C(r; \boldsymbol{\xi})$ be a covariance function parametrised by the m -dimensional parameter vector $\boldsymbol{\xi}$. The elements of $\boldsymbol{\xi}$ are modelled via

$$g_j(\xi_j) = \mathbf{A}_{j,t}\boldsymbol{\beta}_j + \sum_i f_{j,i}(x_t^{S_{j,i}}), \quad \text{for } j = 1, \dots, m, \quad (6)$$

a Generalised Additive Model, the parameters of which (including regularisation) are to be estimated. Details in [Browell et al., 2022].

In the MCD case, we model a subset of the elements of $\boldsymbol{\eta}$ in exactly the same way, selection precisely of which η_i to model is made via a boosting algorithm. Details in [Gioia et al., 2022].

EXAMPLES

The only difference between models is the covariance structure, all margins/density forecasts are the same.

We use standard scoring rules for multivariate probabilistic forecasting:

- Multivariate Energy Score (generalisation of CRPS)
- Log (or Ignorance) Score
- Variogram Score (with $p = 0.5$ and $p = 1$)

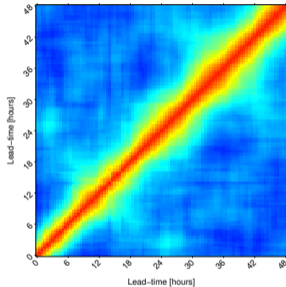


Figure 3: Empirical temporal dependency structure of wind power forecasts

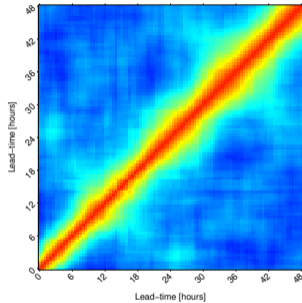
The temporal dependency structure of wind power forecast is non-stationary and complex.

Modelled with exponential correlation function and cubic splines: θ becomes a smooth function of lead-time

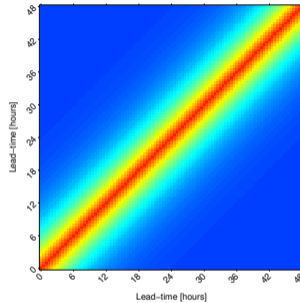
$$\theta = \hat{\theta}_{cr}(d) = \beta_0 + f_{cr}(d) \quad .$$

where d is distance along the diagonal.

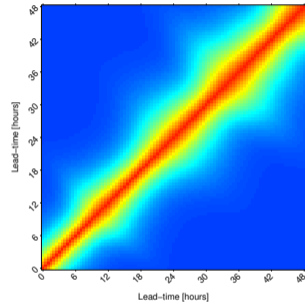
TEMPORAL STRUCTURE IN WIND POWER FORECASTS



(a) Empirical



(b) Constant $\hat{\theta}$ (Stationary)



(c) GAC $\hat{\theta}_{cr}(d)$

Figure 4: Temporal dependency structure of wind power forecasts from 0 to 48 hours-ahead. Forecasts have a visible non-stationary structure. The width of the diagonal ridge indicates how long forecast errors are likely to persist for in time.

Table 2: Results different temporal dependency models for wind power forecasting. Underline indicates that the corresponding skill score relative to the GAC model are not significantly different from zero.

Name	Energy	Log	VS-0.5	VS-1
Empirical	<u>7.139</u>	Inf	<u>1409</u>	<u>5444</u>
Constant	<u>7.142</u>	19.86	<u>1409</u>	<u>5439</u>
GAC	7.137	15.46	1406	5433

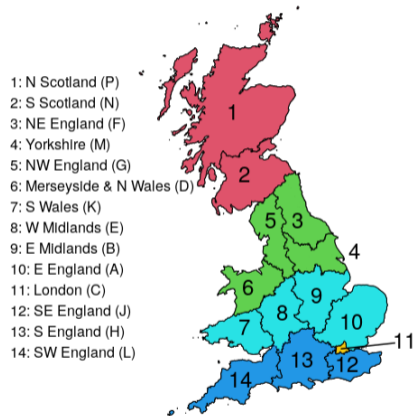
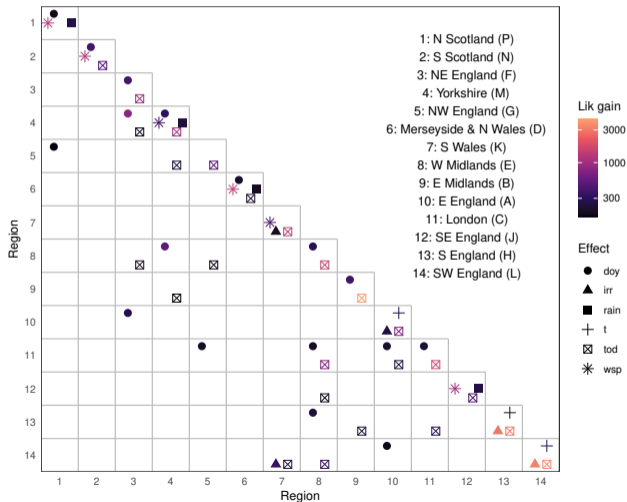


Figure 5: A map of the regions (Grid Supply Point groups) forming GB's electricity grid.

Figure 6: Model selection results. The diagonal corresponds to the elements of D , the rest to those of T .



	Scot - Rest		South - Rest		Lon - Neigh.	
	Log	CRPS	Log	CRPS	Log	CRPS
<i>Indep</i>	5879	6169	4495	4310	2842	1491
<i>Static</i>	4645	5790	4206	4221	2850	1489
<i>Cal</i>	4543	5701	<u>4117</u>	<u>4150</u>	2715	1454
<i>Cal+Ren</i>	<u>4541</u>	<u>5698</u>	4121	<u>4150</u>	<u>2695</u>	<u>1450</u>
<i>Full</i>	4545	5703	4122	4153	2703	1452

Table 3: Day-ahead performance when forecasting the marginal distribution of differences in net-demand across key boundaries. The best score in each column has been underlined.




SUMMARY AND DISCUSSION

- We present two approaches to model **dynamic** and **non-stationary** covariance structures flexibly
- Doing so may **substantially improve** the quality of multi-variate probabilistic forecasts (and other covariance-based models!)
- There is much still to be done to understand and improve model selection and estimation...
- Weather-related uncertainty is inherent in ensemble numerical weather prediction. How can we combine weather and non-weather dependency structures?

Full details including more examples, code and data can be found in [Browell et al., 2022, Gioia et al., 2022].

Papers, slides, code and more linked from www.jethrobrowell.com

References:

-  Browell, J., Gilbert, C., and Fasiolo, M. (2022). Covariance structures for high-dimensional energy forecasting. *Electric Power Systems Research*, 211:108446.
-  Gioia, V., Fasiolo, M., Browell, J., and Bellio, R. (2022). Additive Covariance Matrix Models: Modelling Regional Electricity Net-Demand in Great Britain.
-  Pourahmadi, M. (1999). Joint Mean-Covariance Models with Applications to Longitudinal Data: Unconstrained Parameterisation. *Biometrika*, 86(3):677–690.

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